Chapter - 3: Functions of Random Variables

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Input	Output
$X(t_0) \ge 0.5$	Y=1
$X(t_0) < 0.5$	Y=0

• If Y is a RV, we can write, $\{Y = 0\} = \{X < 0.5\}$ $\{Y = 1\} = \{X \ge 0.5\}$



• If X: N(1,1) then we can calculate the probabilities as



Different Views of Functions of RV

First view $(Y : \Omega \to R_Y)$. For every $\zeta \in \Omega$, we generate a number $g[X(\zeta)] \triangleq Y(\zeta)$. The rule Y, which generates the numbers $\{Y(\zeta)\}$ for random outcomes $\{\zeta \in \Omega\}$, is a r.v. with domain Ω and range $R_Y \subset R$. Finally for every Borel set of real numbers B_Y , the set $\{\zeta : Y(\zeta) \in B_Y\}$ is an event. In particular the event $\{\zeta : Y(\zeta) \leq y\}$ is equal to the event $\{\zeta : g[X(\zeta)] \leq y\}$. In this view, the stress is on Y as a mapping from Ω to R_Y . The intermediate role of

X is suppressed.



View #2

Second view (input/output systems view). For every value of $X(\zeta)$ in the range R_X we generate a new number Y = g(X) whose range is R_Y . The rule Y whose domain is R_X and range is R_Y is a function of the random variable X. In this view the stress is on viewing Y as a mapping from one set of real numbers to another. A model for this view is to regard X as the input to a system with transmittance $g(\cdot)$. For such a system, an input x gets transformed to an output y = g(x) and an input function X gets transformed to an output function Y = g(X). (See Figure 3.1 - 4.)

$$egin{aligned} \{\zeta:Y(\zeta)\leq y\}&=\{\zeta:g[X(\zeta)]\leq y\}\ &=\{\zeta:X(\zeta)\in C_y\} \end{aligned}$$



Example 3.2-1

• Let **Y** = 2X+3 and X: U(0,1). Find $F_Y(y)$

• Solution: $\{Y \le y\} = \{2X + 3 \le y\} = \{X \le \frac{1}{2}(y - 3)\}$

Hence C_y is the interval $\left(-\infty, \frac{1}{2}(y-3)\right)$ and $F_Y(y) = F_X\left(\frac{y-3}{2}\right)$

The pdf of Y is

$$egin{aligned} f_Y(y) &= rac{dF_Y(y)}{dy} = rac{d}{dy}iggl[F_X\left(rac{y-3}{2}
ight)iggr] \ &= rac{1}{2}f_X\left(rac{y-3}{2}
ight) \ \hline \end{aligned}$$

• <u>Generalization</u>

• NOTE: Read the text before ex 3.2-2 on discrete r.v.

$$\{Y\leq y\}=\{aX+b\leq y\}=\left\{X\leq rac{y-b}{a}
ight\}$$

Since **a** is negative, the inequality is flipped. Whenever you multiply or divide an inequality by a negative number, you must flip the inequality sign.

▶ But for
$$a < 0, \{Y \le y\} = \{aX + b \le y\} = \{aX \le y - b\} = \left\{X \ge \frac{y - b}{a}\right\} = \{X \ge -A\}$$

$$F_Y(y) = P[X \ge -A] = P[X \ge -A] = 1 - P[X \le -A] = 1 - F_X(-A)$$

This is only true for
continuous r.y.



Example for a<0

$$Y = 2X+3$$
, $a = 2$, $b = 3$



$$F_Y(y) = P[\{Y \le 10\}] = P[\{X \le 7/2\}] = F_X(7/2)$$



 $egin{aligned} F_Y(y) &= P[\{Y \leq 10\}] = P[\{-2X \leq y-3\}] \ &= P[\{X \geq (10-3)/(-2)\}] = P[\{X \geq -7/2\}] \end{aligned}$

$$P[\{X \ge -7/2\}] = P[\{X \ge -7/2\}] = 1 - P[\{X \le -7/2\}] = 1 - F_X(-7/2)$$

This is only true for continuous r.v.

because P[X=7/2] does not exist

Example 3.2-5



(Transformation of PDF's.) Let X have a continuous PDF $F_X(x)$, which is a strict monotone increasing function[†] of x. Let Y be an r.v. formed from X by the transformation

$$Y = F_X(X).$$
 (3.2-12)

To compute $F_Y(y)$, we proceed as usual:

$$\{Y \le y\} = \{F_X(X) \le y\}$$

= $\{X \le F_X^{-1}(y)\}.$

Hence

$$F_Y(y) = P[F_X(X) \le y]$$

= $P[X \le F_X^{-1}(y)]$
= $\int_{\{x:F_X(x) \le y\}} f_X(x) dx$

- 1. Let y < 0. Then since $0 \le F_X(x) \le 1$ for all $x \in [-\infty, \infty]$, the set $\{x: F_X(x) \le y\} =$ ϕ and $F_Y(y) = 0$. 2. Let y > 1. Then $\{x: F_X(x) \le y\} = [-\infty, \infty]$ and $F_Y(y) = 1$.
- 3. Let $0 \le y \le 1$. Then $\{x: F_X(x) \le y\} = \{x: x \le F_X^{-1}(y)\}$

and

$$F_Y(y) = \int_{-\infty}^{F_X^{-1}(y)} f_X(x) dx = F_X(F_X^{-1}(y)) = y.$$

$$F_Y(y) = \begin{cases} 0, & y < 0\\ y, & 0 \le y \le 1\\ 1, & y > 1. \end{cases}$$
(3.2-13)

Equation 3.2-13 says that whatever probability law X obeys, $Y \triangleq F_X(X)$ will be a uniform r.v. Conversely, given a uniform r.v. Y, the transformation $X \stackrel{\Delta}{=} F_X^{-1}(Y)$ will generate a r.v. with PDF $F_X(x)$ (Figure 3.2-5). This technique is sometimes used in *simulation* to generate r.v.'s with specified distributions from a uniform r.v.

Example 3.2-2 - Multiple roots of Y = g(X)

- **Define** $Y = X^2$
- We can write the set of outcomes for the RV Y

$$\{Y \le y\} = \{X^2 \le y\} = \{-\sqrt{y} \le X \le \sqrt{y}\}$$
$$= \{-\sqrt{y} < X \le \sqrt{y}\} \cup \{X = -\sqrt{y}\}$$
$$\mathsf{CDF} \twoheadrightarrow F_Y(y) = F_X(\sqrt{y}) - F_X(-\sqrt{y}) + P[X = -\sqrt{y}]$$
$$\mathsf{PDF} \twoheadrightarrow f_Y(y) = \frac{d}{du}[F_Y(y)]$$

(Continuous RV)

$$egin{aligned} & f_Y(y) = rac{1}{dy} [F_Y(y)] \ & = rac{1}{2\sqrt{y}} f_X(\sqrt{y}) + rac{1}{2\sqrt{y}} f_X(-\sqrt{y}) \end{aligned}$$

• If X: N(0,1)

$$f_Y(y) = rac{1}{\sqrt{2\pi y}} e^{-rac{1}{2}y} u(y)$$
 - WHY?

Example 3.2-8 - Multiple roots of Y = g(X)

for $0 \le y \le 1$, the event $\{Y \le y\}$ satisfies

$$\{Y \le y\} = \{\sin X \le y\} \\ = \{\pi - \sin^{-1} y < X \le \pi\} \cup \{-\pi < X \le \sin^{-1} y\}.$$

Since the two events on the last line are disjoint, we obtain

$$F_Y(y) = F_X(\pi) - F_X(\pi - \sin^{-1} y) + F_X(\sin^{-1} y) - F_X(-\pi).$$

Hence

$$f_Y(y) = \frac{dF_Y(y)}{dy} = f_X(\pi - \sin^{-1}y)\frac{1}{\sqrt{1 - y^2}}$$
$$+f_X(\sin^{-1}y)\frac{1}{\sqrt{1 - y^2}}$$
$$= \frac{1}{\pi}\frac{1}{\sqrt{1 - y^2}} \qquad 0 \le y < 1.$$





The red line defines the set $Y \leq y$

$$P[\{-\pi - sin^{-1} < X \le sin^{-1}\}] = F_X(sin^{-1}y) - F_X(-\pi - sin^{-1}y) \longrightarrow f_Y(y) = rac{1}{\pi}rac{1}{\sqrt{1-y^2}} \quad ext{ for } -1 < y < 0$$

Generalization for roots of Y=g(X)



•
$$\int_{i=1}^{n} f_X(x_i) |dx_i|$$
•
$$\int_{i=1}^{n} f_X(x_i) \left| \frac{dx_i}{dy} \right| = \sum_{i=1}^{n} f_X(x_i) \left| \frac{dy}{dx_i} \right|^{-1} = \sum_{i=1}^{n} f_X(x_i) / |g'(x_i)|$$
Note the

Note the modulus for $g^{\prime}(x)$

Example 3.2-9

• The roots in example 3.2-8 were, for y > 0 are $x_1 = \sin^{-1} y, x_2 = \pi - \sin^{-1} y$

At
$$x_1 = \sin^{-1} y$$
 we get $\left. dg/dx \right|_{x=x_1} = \cos(\sin^{-1} y)$

$$\frac{dg}{dx}\Big|_{x=x_2} = \cos\left(\pi - \sin^{-1}y\right) = \cos\pi\cos\left(\sin^{-1}y\right) + \sin\pi\sin\left(\sin^{-1}y\right)$$
$$= -\cos\left(\sin^{-1}y\right)$$



•
$$\left| \frac{dg}{dx} \right|_{x_1} = \left| \frac{dg}{dx} \right|_{x_2} = \sqrt{1 - y^2}$$

$$f_Y(y)=rac{1}{\pi}rac{1}{\sqrt{1-y^2}}\quad 0\leq y<1$$

Practice - Find pdf of **Y** = **cos(X)** using the same principle?

Example 3.2-11

• We know
$$f_{Y}(y) = \sum_{i=1}^{n} f_{X}(x_{i}) / |g'(x_{i})|$$

• Find roots of **y=g(x)**

 \circ Thus, g'(x) = 0 for $|x| \geq 1,$ and g'(x) = 1 for -1 < x < 1

 \circ For $y \ge 1$ and $y \le -1$ there are no real roots to $y - g(x) = 0 \longrightarrow f_Y(y) = 0$ g(x) = x, -1 < x < 0

$$\circ$$
 For $-1 < y < 1$, $|g'(x)| = 1 \longrightarrow f_Y(y) = f_X(y)$

• If $X : \mathcal{N}(0,1)$

$$f_Y(y) = \begin{cases} 0, & |y| \ge 1\\ (2\pi)^{-1/2} \exp\left\{-\frac{1}{2}y^2\right\} + 0.317\delta(y), & -1 < y < 1 \end{cases}$$

$$P[Y = 0] = P[X \ge 1] + P[X \le -1]$$

$$= (1 - P[X \le 1]) + P[X \le -1]]$$

$$= (1 - \{1/2 + erf(1)\}) + (1/2 + erf(-1))$$

$$= 1/2 - erf(1) + 1/2 - erf(1)$$

$$= 1 - 2 \times 0.34134 \text{ (From Table 2.4-1 in text)}$$

Justification: This adjustment is required because when g'(x) = 0 it is a flat region in y=g(x). So, for any x in that flat region the yvalues are identical. This will create a probability mass at that value of y. The mass is equal to the probability of the event X falls in the flat area. In this case $X \ge 1$ and $X \le -1$



Infinite roots - Example 3.2-12

• The excursions of y = g(x) suggests the same roots as before for |y| > 1

• But infinite roots for
$$-1 < y < 1$$

 $\circ x_n = y + 2n$ With $|g'(x)| = 1$ at each root

- So, $f_Y(y) = \sum_{n=-\infty}^{\infty} f_X(y+2n) \operatorname{rect}\left[\frac{y}{2}\right]$
- If $X : \mathcal{N}(0, 1)$, integrate to check if equal to 1

$$egin{aligned} &\int_{-\infty}^\infty f_Y(y) dy = rac{1}{\sqrt{2\pi}} \sum_{n=-\infty}^\infty \int_{-1}^1 \expigg\{ -rac{1}{2} (y+2n)^2 igg\} dy \ &= rac{1}{\sqrt{2\pi}} \sum_{n=-\infty}^\infty \int_{-1+2n}^{1+2n} \expigg\{ -rac{1}{2} y^2 igg\} dy \ &= \sum_{n=-\infty}^\infty \{ \operatorname{erf}(1+2n) - \operatorname{erf}(-1+2n) \} \end{aligned}$$

• All terms cancels, except the *second term of the first sum* and *the first term of the last sum*

$$\int_{-\infty}^{\infty} f_Y(y) dy = -\operatorname{erf}(-\infty) + \operatorname{erf}(\infty) = 2 imes \operatorname{erf}(\infty) = 1$$



Two variable functions Z = g(X,Y)

• Definition

 $\{\zeta: Z(\zeta) \leq z\} ext{ and } \{\zeta: X(\zeta), Y(\zeta) \in C_z\}$

• Using shorter notations

 $\{Z\leq z\}=\{(X,Y)\in C_z\}$ $F_Z(z)=\iint_{(x,y)\in C_z}f_{XY}(x,y)dxdy$

- Other functions include
 - $\circ \quad \mathsf{Z} = \max(\mathsf{X},\mathsf{Y}),$
 - \circ Z = X+Y, aX+bY
 - $\circ \qquad \mathsf{Z}=\mathsf{X}^2{+}\mathsf{Y}^2$
 - o $Z = (X^2 + Y^2)^{1/2}$





Example 3.3-1: Z=XY

- For z>0 $F_Z(z) = \int_0^\infty \left(\int_{-\infty}^{z/y} f_{XY}(x,y) dx \right) dy + \int_{-\infty}^0 \left(\int_{z/y}^\infty f_{XY}(x,y) dx \right) dy$
- The CDF is given by

$$F_{Z}(z) = \int_{0}^{\infty} [G_{XY}(z/y, y) - G_{XY}(-\infty, y)] dy \ + \int_{-\infty}^{0} [G_{XY}(\infty, y) - G_{XY}(z/y, y)] dy \ f_{Z(z)} = rac{dF_{Z}(z)}{dz} = \int_{-\infty}^{\infty} rac{1}{|y|} f_{XY}(z/y, y) dy$$



• Calculate $f_Z(z)$ if X and Y are iid Cauchy variables,

$$f_X(x)=f_Y(x)=rac{lpha/\pi}{lpha^2+x^2}$$

Note: the limits on the second term for the integral over \mathbf{x} , is also from a smaller (-ve) number \mathbf{z}/\mathbf{y} to $+\infty$

$$y = \frac{z}{x}$$

Sum of two variable: Z = X+Y

• Proceeding as per definition

$$egin{aligned} F_Z(z) &= \iint_{x+y \leq z} f_{XY}(x,y) dx dy \ &= \int_{-\infty}^\infty \left(\int_{-\infty}^{z-y} f_{XY}(x,y) dx
ight) dy \ &= \int_{-\infty}^\infty \left[G_{XY}(z-y,y) - G_{XY}(-\infty,y)
ight] dy \ &= \int_{-\infty}^\infty \left[G_{XY}(x,y) dx
ight] dy \ &= \int_{-\infty}^\infty \left[G_{XY}(x,y) d$$



$$f_{Z}(z) = \frac{dF_{Z}(z)}{dz} = \int_{-\infty}^{\infty} \frac{d}{dz} [G_{XY}(z - y, y)] dy$$

=
$$\int_{-\infty}^{\infty} f_{XY}(z - y, y) dy = \int_{-\infty}^{\infty} f_{X}(z - y) f_{Y}(y) dy \text{ (if X and Y are independent)}$$

This is called the convolution equation. Flipped image of $f_{x}(y)$ and shifted to the left by z units gives, $f_{x}(-y+z)$

y

y = z - x

x

Example 3.3-4

• X and Y are r.v. with

 $f_X(x)=e^{-x}u(x)$

$$f_Y(y) = rac{1}{2}[u(y+1) - u(y-1)] \qquad Z = X + Y$$



 $f_X(z-y)=e^{-(z-y)}u(z-y)$

Sanity check: $f_z(z)$ is continuous since there are no delta functions involved in the integration. Check with z = 1 in region 1 and region 2 to confirm equality, which should be the case for continuous functions.

What about Z = aX+bY

• Let a>0, b>0,
$$y = \frac{z}{b} - \frac{ax}{b}$$

$$egin{aligned} F_Z(z) &= \iint_{g(x,y) \leq z} f_{XY}(x,y) dx dy \ &= \int_{-\infty}^\infty f_Y(y) \left(\int_{-\infty}^{z/a - by/a} f_X(x) dx
ight) dy \end{aligned}$$

Differentiating, w.r.t z

$$f_Z(z) = rac{1}{a} \int_{-\infty}^\infty f_X\left(rac{z}{a} - rac{by}{a}
ight) f_Y(y) dy$$

• Another way to solve this is to define new r.v.,

V = aX, W = bY

Then apply Z = V + W and convolve as before.

Substitute *w* = *by* to get the integral as above

 $1 \circ (\alpha)$



$Z = X^2 + Y^2$ and $Z = (X^2 + Y^2)^{1/2}$

• If X,Y are iid with $\mathcal{N}(0,\sigma^2)$, proceed as before,

$$egin{aligned} F_Z(z) &= \iint_{(x,y)\in C_z} f_{XY}(x,y) dx dy \quad ext{for} \quad z \geq 0 \ &= rac{1}{2\pi\sigma^2} \iint_{x^2+y^2 \leq z} e^{-(1/2\sigma^2)\,(x^2+y^2)}\, dx dy \end{aligned}$$

• Converting to polar coordinates

$$egin{aligned} x = r\cos heta & y = r\sin heta \ dxdy &
ightarrow rdrd heta \end{aligned}$$

Then $x^2 + y^2 \leq z \rightarrow r \leq \sqrt{z}$ and the above equation becomes $F_Z(z) = rac{1}{2\pi\sigma^2} \int_0^{2\pi} d heta \int_0^{\sqrt{z}} r \exp\left[-rac{1}{2\sigma^2}r^2
ight] dr = \left[1 - e^{-z/2\sigma^2}
ight] u(z)$

• Differentiating w.r.t **z**

$$f_Z(z)=rac{dF_Z(z)}{dz}=rac{1}{2\sigma^2}e^{-z/2\sigma^2}u(z)$$

Multiple functions of R.V.s: V=g(X,Y), W = h(X,Y)

• **Problem:** Compute joint distribution $F_{VW}(v, w)$ from $F_{XY}(x, y)$

$$egin{aligned} P[V \leq v, W \leq w] &= F_{VW}(v,w) \ &= \iint_{(x,y) \in C_{vw}} f_{XY}(x,y) dx dy \end{aligned}$$

The region C_{vw} is given by the points x, y that satisfy $C_{vw} = \{(x,y) : g(x,y) \le v, h(x,y) \le w\}$

To integrate, express *x* and *y* in terms of *v* and *w*. (see example 3.4-1)

$$F_{VW}(v,w) = \int_{-\infty}^{(v+w)/2} \left(\int_{x-w}^{v-x} f_{XY}(x,y) dy\right) dx$$

$$f_{VW}(v,w) = \frac{\partial^2 F_{VW}(v,w)}{\partial v \partial w}$$
Study the integration from the text

 $y = \frac{v - w}{2}$

Lower limit for the integration

over v = x - w

x + y = v

Upper limit for

 $x = \frac{v + w}{2}$

Simpler Approach

- In the infinitesimal small region, $\{v < V \le v + dv, w < W \le w + dw\}$
- Now, $P[\{v < V \le v + dv, w < W \le w + dw\}]$ is the probability that V and W lie in the infinitesimal rectangle of area $\partial v \partial w$.



..contd

Recall in single variable case in chapter 2 $P[\{x < X \le x + \Delta x\}] \simeq f_X(x). \Delta x$

• Therefore, we can write,

$$egin{aligned} P[v < V &\leq v + dv, w < W \leq w + dw] = \iint_{\mathscr{R}} f_{VW}(\xi,\eta) d\xi d\eta \ &= f_{VW}(v,w) A(\mathscr{R}) \ &= \iint_{\mathscr{S}} f_{XY}(\xi,\eta) d\xi d\eta \ &= f_{XY}(x,y) A(\mathscr{S}) \end{aligned}$$

 $egin{aligned} ext{Note:} & P(B) = \iint_{\mathscr{R}} f_{XY}(x,y) dx dy
eq \iint_{\mathscr{R}} f_{XY}(\phi(v,w),arphi(v,w)) dv dw \ ext{\rightarrow} & ext{Because the area (or volume), } dx dy
eq dv dw \end{aligned}$

- The area of the parallelogram $P_1P_2P_3P_4$ in vector notation is given by $A(\mathscr{S}) = \left| \overrightarrow{P_1P_2} \times \overrightarrow{P_1P_3} \right| = \left| \left(\frac{\partial \phi}{\partial v} dv \mathbf{i} + \frac{\partial \varphi}{\partial v} \mathbf{j} dv \right) \times \left(\frac{\partial \phi}{\partial w} \mathbf{i} dw + \frac{\partial \varphi}{\partial w} \mathbf{j} dw \right) \right|$ $= \left| \frac{\partial \phi}{\partial v} \frac{\partial \varphi}{\partial w} - \frac{\partial \phi}{\partial w} \frac{\partial \varphi}{\partial v} \right| dv dw$
- Therefore, the Jacobian of functions $\mathbf{x} = \boldsymbol{\phi}(\mathbf{v}, \mathbf{w})$ and $\mathbf{y} = \boldsymbol{\psi}(\mathbf{v}, \mathbf{w})$ is $|\tilde{J}_i| = \max \begin{vmatrix} \partial \phi_i / \partial v & \partial \phi_i / \partial w \\ \partial \varphi_i / \partial v & \partial \varphi_i / \partial w \end{vmatrix} = |\partial \phi_i / \partial v \times \partial \varphi_i / \partial w - \partial \varphi_i / \partial v \times \partial \phi_i / \partial w|$

.. contd



$$J = egin{bmatrix} rac{\partial g}{\partial x} & rac{\partial g}{\partial y} \ rac{\partial h}{\partial x} & rac{\partial h}{\partial y} \end{bmatrix} = { ilde J}_i^{-1}$$

Example 3.5-2

We are given two functions

$$v \stackrel{\Delta}{=} g(x, y) = 3x + 5y$$

 $w \stackrel{\Delta}{=} h(x, y) = x + 2y$

and the joint pdf f_{XY} of two r.v.'s X, Y. What is the joint pdf of two new random variables V = g(X, Y), W = h(X, Y)?

Solution The inverse mappings are computed from Equation 3.4-13 to be

$$x = \phi(v, w) = 2v - 5w$$
$$y = \Phi(v, w) = -v + 3w.$$

Then

$$\frac{\partial \phi}{\partial v} = 2, \frac{\partial \phi}{\partial w} = -5, \frac{\partial \Phi}{\partial v} = -1, \frac{\partial \Phi}{\partial w} = 3$$

and

$$|\tilde{J}| = \max \begin{vmatrix} 2 & -5 \\ -1 & 3 \end{vmatrix} = 1.$$

Assume $f_{XY}(x, y) = (2\pi)^{-1} \exp[-\frac{1}{2}(x^2 + y^2)]$. Then, from Equation 3.4-11

$$f_{VW}(v,w) = \frac{1}{2\pi} \exp\left[-\frac{1}{2}[(2v-5w)^2 + (-v+3w)^2]\right]$$
$$= \frac{1}{2\pi} \exp\left[-\frac{1}{2}(5v^2 - 26vw + 34w^2)\right].$$

Example 3.5-5

